THE RELATIONSHIP BETWEEN SENSITIVITY AND VARIABILITY IN NORMAL AND GLAUCOMATOUS VISUAL FIELDS

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Abstract

Purpose: To establish the relationship between sensitivity and variability in normal, ocular hypertensive and glaucomatous eyes.

Method: Frequency-of-seeing (FOS) data were collected from four visual field locations of 64 eyes (22 normal, six ocular hypertensive and 38 glaucomatous), using a constant stimulus method on a Henson 4000 perimeter. At each location, 20 stimuli (0.5°) were presented (200 msec) at each of six intensities straddling the estimated threshold. In glaucomatous eyes, at least one location was chosen to lie in an area of normal sensitivity. The FOS data were fitted with a cumulative normal distribution, the standard deviation (SD) of which was used as an estimate of response variability.

Results: Variability was found to increase with decreased sensitivity. The relationship was best described by the function \( \log_e(SD) = a \times \text{sensitivity(dB)} + b \), where the constants \( a \) and \( b \) were -0.08 and 3.22, respectively.

Conclusions: Previously presented linear models of the relationship between sensitivity and variability fail to accurately predict variability at high and low perimetric sensitivities, suggesting negative variability at high sensitivities and underestimating variability at low dB values. Our model attempts to correct these shortcomings. It can be used in computer simulations of perimetry to estimate the performance of both existing and new perimetric algorithms.

Background

An important parameter in the development and assessment of perimetric algorithms is the relationship between sensitivity and variability. Variability has been known for some time to increase when sensitivity reduces. The magnitude of the effect can be very large; in areas of moderate sensitivity loss (8-18dB), the 95% confidence intervals can exceed the measurement range of the Humphrey Visual Field Analyzer.

Early research used repeat measures of the threshold to quantify variability (fluctuation), but more recent research has concentrated upon the use of frequency-of-seeing (FOS) curves. The slope of the FOS curve gives a measure of response variability, while the 50% seen level gives a measure of sensitivity. FOS curves fully define the

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stimulus response relationship and can be used in simulations to predict the performance of perimetric test algorithms.

The relationship between FOS-derived measures of sensitivity and response variability has been reported by Weber and Rau\textsuperscript{9}, Olsson \textit{et al.}\textsuperscript{10} and Chauhan \textit{et al.}\textsuperscript{11}. Weber and Rau\textsuperscript{9} measured variability in normal eyes (central and peripheral locations) and in glaucomatous eyes. They found that a straight line could fit all their results with normal central locations at one end and defective glaucomatous locations at the other. Chauhan \textit{et al.}\textsuperscript{11}, while demonstrating a similar relationship, emphasized the poor quality of the linear fit at lower sensitivities. Olsson \textit{et al.}\textsuperscript{10} found a good relationship between $\log_e(\text{variability})$ and threshold, pointing out that this transformation of the variability scale made the residuals reasonably independent of the variability.

We re-visited the relationship between sensitivity and variability and concluded that the relationship is best expressed by a model where $\log(\text{variability})$ is a linear function of sensitivity.

\textbf{Method}

\textit{Subjects}

Data were collected from one eye of 22 normal subjects (median age 33, range 19-71 years), six ocular hypertensives (OHT) (median age 69.4, range 66-83 years), and 36 patients with primary open-angle glaucoma (POAG) (median age 67.5, range 33 to 80 years). Normal subjects were recruited from hospital staff, had no history of ophthalmic disease, a normal ophthalmic examination, no systemic illness, and a visual acuity better than 6/9. The OHTs were recruited from the glaucoma clinics at the Manchester Royal Eye Hospital. Inclusion/exclusion criteria were the same as for the normal subjects with the exception of a raised IOP. The POAG patients were also recruited from the glaucoma clinics at Manchester Royal Eye Hospital. They all had a confirmed diagnosis of POAG with visual field loss, stable controlled IOP, no other ocular pathology, previous experience with automated perimetry, and a visual acuity of better than 6/18.

\textit{Test locations}

FOS data were collected at four visual field locations during a single experimental session. For the normal and ocular hypertensive subjects, the locations were 12.7° from fixation along the 45, 135, 225 and 315 meridians. For the glaucoma patients, one location was chosen to be in an area of normal sensitivity, and, if possible, three locations in or adjacent to a damaged area of the visual field. The damaged locations were chosen on the basis of a Humphrey Visual Field Analyzer (Program 24-2) test performed prior to the collection of FOS data.

\textit{Data collection}

A computer program was specifically written to allow the collection of FOS data on a Henson 4000 bowl perimeter\textsuperscript{14}. All stimuli subtended 0.5° and were presented for 200
msec with a maximum luminance increment of 1000 cd/m². After specification of the test locations, the program estimated the sensitivity at each location using the standard full threshold (4-2) strategy. For each location the program then selected five intensities which straddled the estimated threshold in 2dB steps. During each session, the experimenter, who received continuous feedback on the selected intensities and current responses, would repeatedly adjust the intensities and number of presentations to ensure that 1. the response range approached 0 and 100% seen; 2. data were collected for at least six intensities; and 3. there were a minimum of 20 presentations at each intensity. While the adjustments were being made, the subject was allowed a short rest. The presentation of stimuli was randomized with the inclusion of false positive and false negative response trials.

Data analysis

The FOS data from each test location were imported into the statistical package SPSS for probit regression analysis. The standard deviation (SD) of the fitted cumulative normal function was used as an estimate of response variability.

Results and discussion

Figure 1 gives typical FOS data from one location in a normal eye (a) and two locations in a glaucomatous eye (b and c), one of which is from an area of reduced sensitivity (c). It can be seen that, while the data from the glaucomatous eye’s location with normal sensitivity are similar to that from the normal eye, the location with reduced sensitivity has a much shallower curve.

When deciding upon appropriate test locations for the POAG eyes, one location was chosen to lie in an area of near normal sensitivity while the other three were chosen to lie in areas of reduced sensitivity. Many of the reduced sensitivity locations were subsequently found to yield results that, due to limitations in the dynamic range of the perimeter, did not approach the 100% seen level and could not, therefore, be accurately fitted by the probit function. These locations have not been included in the subsequent analysis.

As sensitivity decreases, there is an increase in response variability (Fig. 2). This finding is in agreement with the earlier work of Weber and Rau⁵, Olsson et al.¹⁰ and Chauhan et al.¹¹. The data also show an increased scatter of SD values with reduced sensitivity, a finding that was also noted by Chauhan et al.¹¹.

A linear fit to the data of Figure 2 underestimates variability at high and low sensitivities. It also predicts negative variability at high sensitivity, which is clearly at odds with common sense. The data in Figure 2 also show inequality of variance that violates the assumptions of the least-squares method.

Figure 3 gives the same data after a log transform of the SD values as proposed by Olsson et al.¹⁰. The solid line (least-squares linear regression) now accurately represents the data. The parameters of the model \( \log_e(SD) = a \times \text{sensitivity} + b \) were -0.08 and 3.22, respectively. Parameter b is dependent upon the decibel scale of the instrument. This model no longer predicts negative response variability at high sensitivities. The data points are now normally distributed around the regression line with equality of variance. The equation of the best linear fit to these data has been transformed and superimposed upon the data in Figure 2.

The relationship between sensitivity and variability while being similar to that of Weber
Fig. 1. Three frequency-of-seeing curves. a. Normal eye; b. glaucomatous eye normal sensitivity; and c. glaucomatous eye reduced sensitivity.
and Rau⁸ is much tighter than that reported by Chauhan et al.¹¹. This can, in part, be explained by differences in experimental methods. Weber and Rau⁸ and the present research presented a fairly large number of stimuli at each intensity (25 and 20), while Chauhan et al.¹¹ only presented five. With only five presentations, the precision of any FOS estimate is low, and even single response errors can result in large differences of the fitted function.

Another difference in the methodology is that Weber and Rau⁸ and Chauhan et al.¹¹ collected data at a greater number of stimulus intensities (ten and 17 versus a minimum of six in the present research). A large number of stimulus intensities leads to an increase in test time and invariably results in several test intensities with response probabilities close

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**Fig. 2.** Variability versus sensitivity linear axes.

**Fig. 3.** Variability versus sensitivity.

\[ SD = 25 \times \exp(-0.08 \times \text{sensitivity}) \]

\[ \log_{10}(SD) = -0.08 \times \text{sensitivity} + 3.22 \]
to zero or one, which contribute little to the estimate of the FOS curve. For maximum information on the gradient of the FOS curve, data need to be collected at the shoulder areas of the curve. The technique used in this study allowed the operator to adjust the number of presentations and the intensities during the test. This ensured that most of the data were collected at intensities with high informational value and, thereby, contributed more to the accuracy of the FOS estimates. The more efficient approach adopted in the present research also meant that FOS curves could be determined from four visual field locations with a total number of presentations similar to that of a 30-2 full threshold test. As one of the intentions of the present research was to derive a relationship between sensitivity and response variability which could be used in perimetric simulations, it was felt important that as many parameters as possible, in particular test time, should be similar to those of a standard perimetric test.

The data from normal, OHT and POAG subjects in Figures 2 or 3 are well described by a single model. There is no evidence of increased response variability for visual field locations of normal sensitivity in glaucomatous subjects. Our normal subjects with very high sensitivities showed an increase in their response variability. This might be due to frequent false-positive response errors that would lead to both an overestimation of sensitivity and to higher response variability.

The present research has demonstrated that the relationship between sensitivity and log response variability can be well expressed by a linear function for normal, OHT and POAG eyes. This finding will be useful in computer simulations of perimetry to estimate the performance of both existing and new perimetric algorithms15.

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References

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